

# Some Processes Associated with Fractional Bessel Processes

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## Abstract

Let  $B = \{(B_t^1, \dots, B_t^d), t \geq 0\}$  be a  $d$ -dimensional fractional Brownian motion with Hurst parameter  $H$  and let  $R_t = \sqrt{(B_t^1)^2 + \dots + (B_t^d)^2}$  be the fractional Bessel process. Itô's formula for the fractional Brownian motion leads to the equation  $R_t = \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i + H(d-1) \int_0^t \frac{s^{2H-1}}{R_s} ds$ . In the Brownian motion case ( $H = 1/2$ ),  $X_t = \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i$  is a Brownian motion. In this paper it is shown that  $X_t$  is not a fractional Brownian motion if  $H \neq 1/2$ . We will study some other properties of this stochastic process as well.

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# 1 Introduction

Let  $B = \{(B_t^1, \dots, B_t^d), t \geq 0\}$  be a  $d$ -dimensional fractional Brownian motion with Hurst parameter  $H \in (0, 1)$ . That is, the components of  $B$  are independent one-dimensional fractional Brownian motions with Hurst parameter  $H \in (0, 1)$ .

Denote the fractional Bessel process by  $R_t = \sqrt{(B_t^1)^2 + \dots + (B_t^d)^2}$ . In the standard Brownian motion case there is an extensive literature on this process, see for example [7]. It is natural and interesting to study this process for any other parameter  $H$ . If  $d \geq 2$  and  $1/2 < H < 1$ , using the Itô's formula for the fractional Brownian motion we obtain

$$R_t = \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i + H(d-1) \int_0^t \frac{s^{2H-1}}{R_s} ds \quad (1)$$

and for  $d = 1$  we have

$$|B_t| = \int_0^t \text{sign}(B_s) dB_s + H \int_0^t \delta_0(B_s) s^{2H-1} ds, \quad (2)$$

where  $\delta_0$  is the Dirac delta function, and the stochastic integrals are interpreted in the divergence form. Equation (1) have been proved in [4] in the case  $H > \frac{1}{2}$ , and for Equation (2) we refer to [1], [4], [5] and [6].

In the classical Brownian motion case it is well-known from the Lévy's characterization theorem that the first term in the decomposition (1)

$$X_t = \begin{cases} \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i & \text{when } d \geq 2 \\ \int_0^t \text{sign}(B_s) dB_s & \text{when } d = 1 \end{cases}$$

is a classical Brownian motion. It is then natural and interesting to ask whether for any other  $H$ , the process  $X = \{X_t, t \geq 0\}$  is a fractional Brownian motion or not. The difficulty is that there is no characterization as convenient as Lévy's one for general fractional Brownian motion ( $H \neq 1/2$ ). It is then difficult to show whether a stochastic process is a fractional Brownian motion or not. In this paper, we shall prove that if  $H \neq 1/2$ , then  $\{X_t, t \geq 0\}$  is NOT a fractional Brownian motion. Our approach to show this fact is based on the Wiener chaos expansion (see for example [3] and [5]).

It seems to be the natural method to be used here since there is no other powerful tool available.

Although  $\{X_t, t \geq 0\}$  is not a fractional Brownian motion, it enjoys some properties that the fractional Brownian motion has, such as self-similarity and long range dependence ( $H > 2/3$ ). We will study these and some other properties of the process  $X$ .

Section 2 will recall some preliminary results. Section 3 will study the case  $d = 1$ , namely, the process  $\int_0^t \text{sign}(B_t) dB_t$  and Section 4 is devoted to the study of general dimension, ie, the process  $\sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i$ .

## 2 Preliminaries

Let  $B = \{B_t, t \geq 0\}$  be a fractional Brownian motion with Hurst parameter  $H \in (0, 1)$ . That is,  $B$  is a zero mean Gaussian process with the covariance function

$$R_H(t, s) = E(B_t B_s) = \frac{1}{2} (s^{2H} + t^{2H} - |t - s|^{2H}).$$

We denote by  $K_H(t, s)$  the square integrable kernel such that

$$R_H(t, s) = \int_0^{t \wedge s} K_H(t, u) K_H(s, u) du.$$

Fix a time interval  $[0, T]$ , and let  $\mathcal{H}$  be Hilbert space defined as the closure of the set of step functions on  $[0, T]$  with respect to the scalar product

$$\langle \mathbf{1}_{[0,t]}, \mathbf{1}_{[0,s]} \rangle_{\mathcal{H}} = R_H(t, s).$$

The mapping  $\mathbf{1}_{[0,t]} \longrightarrow B_t$  can be extended to an isometry between  $\mathcal{H}$  and the Gaussian space  $H_1(B)$  associated with  $B$ . We will denote this isometry by  $\varphi \longrightarrow B(\varphi)$ .

The operator defined by

$$(K_H^* \mathbf{1}_{[0,t]})(s) = K_H(t, s) \mathbf{1}_{[0,t]}(s).$$

can be extended to a linear isometry between  $\mathcal{H}$  and  $L^2(0, T)$ . This operator can be expressed in terms of fractional operators. More precisely, if  $H > \frac{1}{2}$  we have

$$(K_H^* \varphi)(s) = c_H \Gamma(H - \frac{1}{2}) s^{\frac{1}{2}-H} (I_{T-}^{H-\frac{1}{2}} u^{H-\frac{1}{2}} \varphi(u))(s)$$

and if  $H < \frac{1}{2}$

$$(K_H^* \varphi)(s) = c_H \Gamma(H + \frac{1}{2}) s^{\frac{1}{2}-H} (D_{T-}^{1-H} u^{H-\frac{1}{2}} \varphi(u))(s),$$

where  $c_H = \sqrt{\frac{2H}{(1-2H)\beta(1-2H, H+1/2)}}$ , and for any  $\alpha > 0$  we denote by  $I_{T-}^\alpha$  (resp.  $D_{T-}^\alpha$ ) the fractional integral (resp. derivative) operator given by

$$I_{T-}^\alpha f(t) = \frac{(-1)^{-\alpha}}{\Gamma(\alpha)} \int_t^T (s-t)^{\alpha-1} f(s) ds$$

(resp.

$$D_{T-}^\alpha f(t) = \frac{(-1)^\alpha}{\Gamma(1-\alpha)} \left( \frac{f(t)}{(T-t)^\alpha} + \alpha \int_t^T \frac{f(t) - f(s)}{(s-t)^{\alpha+1}} ds \right).$$

We denote by  $D$  and  $\delta$  the derivative and divergence operators that can be defined in the framework of the Malliavin calculus with respect to the process  $B$ . Let  $\mathbb{D}^{k,p}$ ,  $p > 1$ ,  $k \in \mathbb{R}$ , be the corresponding Sobolev spaces. We recall that the divergence operator  $\delta$  is defined by means of the duality relationship

$$\mathbb{E}(F\delta(u)) = E \langle DF, u \rangle_{\mathcal{H}}, \quad (3)$$

where  $u$  is a random variable in  $L^2(\Omega; \mathcal{H})$ . We say that  $u$  belongs to the domain of the divergence, denoted by  $\text{Dom } \delta$ , if there is a square integrable random variable  $\delta(u)$  such that (3) holds for any  $F \in \mathbb{D}^{1,2}$ .

The domain of the divergence operator is sometimes too small. For instance, in [1] it is proved that the process  $u = B$  belongs to  $L^2(\Omega; \mathcal{H})$  if and only if  $H > \frac{1}{4}$ . On the other hand, in [2] it is proved that for all  $t \geq 0$ , the process  $\text{sign}(B_t)$  belongs to the domain of the divergence when  $H > \frac{1}{3}$ .

Following the approach of [1] it is possible to extend the domain of the divergence operator to processes whose trajectories are not necessarily in the space  $\mathcal{H}$ . Set  $\mathcal{H}_2 = (K_H^*)^{-1} (K_H^{*,a})^{-1} (L^2(0, T))$ , where  $K_H^{*,a}$  denotes the adjoint of the operator  $K_H^*$ . Denote by  $\mathcal{S}_{\mathcal{H}}$  the space of smooth and cylindrical random variables of the form

$$F = f(B(\phi_1), \dots, B(\phi_n)), \quad (4)$$

where  $n \geq 1$ ,  $f \in C_b^\infty(\mathbb{R}^n)$  ( $f$  and all its partial derivatives are bounded), and  $\phi_i \in \mathcal{H}_2$ .

**Definition 1** Let  $u = \{u_t, t \in [0, T]\}$  be a measurable process such that

$$\mathbb{E} \left( \int_0^T u_t^2 dt \right) < \infty.$$

We say that  $u \in \text{Dom}_T^* \delta$  (extended domain of the divergence in  $[0, T]$ ) if there exists a random variable  $\delta(u) \in L^2(\Omega)$  such that for all  $F \in \mathcal{S}_H$  we have

$$\int_0^T \mathbb{E}(u_t K_H^{*,a} K_H^* D_t F) dt = \mathbb{E}(\delta(u) F).$$

In [1] it is proved that for any  $H \in (0, 1)$ , the process  $\text{sign}(B_t)$  belongs to the extended domain of the divergence in any time interval  $[0, T]$  and the following version of Tanaka's formula holds

$$|B_t| = \int_0^t \text{sign}(B_s) dB_s + H \int_0^t \delta_0(B_s) s^{2H-1} ds. \quad (5)$$

(see also [5], [6] for this and a more general formula). In this formula  $L_t^a = H \int_0^t \delta_0(B_s) s^{2H-1} ds$  is the the density of the occupation measure

$$\Gamma \mapsto 2H \int_0^t 1_\Gamma(B_s) s^{2H-1} ds.$$

### 3 The process $\int_0^t \text{sign}(B_s) dB_s$

Define the process  $X = \{X_t, t \geq 0\}$ , by

$$X_t = \int_0^t \text{sign}(B_s) dB_s.$$

In the case of the classical Brownian motion ( $H = \frac{1}{2}$ ), the process  $X$  turns out to be a Brownian motion. We will show first that for any  $H \in (0, 1)$ ,  $X$  is a  $H$ -self-similar process, that is, for all  $a > 0$  the processes  $\{X_{at}, t \geq 0\}$  and  $\{a^H X_t, t \geq 0\}$  have the same law.

**Proposition 2** The process  $X = \{X_t, t \geq 0\}$  is  $H$ -self-similar.

**Proof.** Using the self-similarity property of the fractional Brownian motion and Tanaka's formula (5) yields that for any  $a > 0$

$$\begin{aligned}
X_{at} &= |B_{at}| - H \int_0^{at} \delta_0(B_s) s^{2H-1} ds \\
&= |B_{at}| - H \int_0^t \delta_0(B_{au})(au)^{2H-1} adu \\
&\stackrel{d}{=} a^H |B_t| - a^{2H} H \int_0^t \delta_0(a^H B_u) u^{2H-1} du \\
&= a^H X_t,
\end{aligned}$$

where the symbol  $\stackrel{d}{=}$  means that the distributions of both processes are the same. This completes the proof. ■

Then, it is natural to conjecture that for any  $H$ , the process  $X_t$  is a fractional Brownian motion of Hurst parameter  $H$ . We will see that this is no longer true if  $H \neq \frac{1}{2}$ , although the process  $X_t$  shares some of the properties of the fractional Brownian motion.

Let us first find the Wiener chaos expansion of the process  $\text{sign}(B_t)$ . We will denote by  $I_n$  the multiple Wiener integral with respect to the process  $B$ .

**Lemma 3** *Let  $0 < H < 1$ . We have the following chaos expansion for  $\text{sign}(B_t)$ :*

$$\text{sign}(B_t) = \sum_{k=0}^{\infty} b_{2k+1} I_{2k+1}(1), \quad (6)$$

where

$$b_{2k+1} = \frac{2(-1)^k}{(2k+1)\sqrt{2\pi}t^{(2k+1)H}k!2^k}.$$

**Proof.** Denote by  $p_{\varepsilon}(x) = \frac{1}{\sqrt{2\pi\varepsilon}}e^{-x^2/\varepsilon}$ ,  $x \in \mathbb{R}$ ,  $\varepsilon > 0$ , the heat kernel. The function

$$f_{\varepsilon}(x) = 2 \int_{-\infty}^x p_{\varepsilon}(y) dy - 1$$

converges to  $\text{sign}(x)$  as  $\varepsilon$  tends to zero. Hence,  $f_{\varepsilon}(B_t)$  converges to  $\text{sign}(B_t)$  in  $L^2(\Omega)$  as  $\varepsilon$  tends to zero. The application of Stroock's formula yields

$$f_{\varepsilon}(B_t) = \sum_{n=0}^{\infty} a_n^{\varepsilon}(t) \int_{0 < s_1 < \dots < s_n < t} dB_{s_1} \cdots dB_{s_n}, \quad (7)$$

where

$$\begin{aligned}
a_n^\varepsilon(t) &= \mathbb{E}[D^n(f_\varepsilon(B_t))] = 2\mathbb{E}[p_\varepsilon^{(n-1)}(B_t)] \\
&= 2(-1)^{n-1} \frac{\partial^{n-1}}{\partial y^{n-1}} \mathbb{E}[p_\varepsilon(B_t - y)]|_{y=0} \\
&= 2(-1)^{n-1} p_{\varepsilon+t^{2H}}^{(n-1)}(0).
\end{aligned}$$

Taking the limit of (7) in  $L^2(\Omega)$  as  $\varepsilon$  tends to zero we obtain

$$\text{sign}(B_t) = \sum_{n=0}^{\infty} a_n(t) \int_{0 < s_1 < \dots < s_n < t} dB_{s_1} \cdots dB_{s_n},$$

where  $a_n(t) = \lim_{\varepsilon \downarrow 0} a_n^\varepsilon(t) = 2(-1)^{n-1} p_{t^{2H}}^{(n-1)}(0)$ . As a consequence,  $a_n(t) = 0$  if  $n$  is even and

$$a_n(t) = \frac{2(-1)^k (2k)!}{\sqrt{2\pi} t^{nH} k! 2^k}$$

if  $n = 2k + 1$ . ■

Using Stirling's formula we obtain

$$\begin{aligned}
\mathbb{E}[I_{2k+1}(b_{2k+1})]^2 &= \frac{4(2k+1)! t^{2H+1}}{(2k+1)2\pi t^{2H+1} (k!2^k)^2} \\
&= \frac{4(2k)!}{(2k+1)2\pi (k!2^k)^2} \\
&\simeq Ck^{-3/2},
\end{aligned}$$

and we have proved the following proposition.

**Proposition 4** For any  $0 < H < 1$ , the random variable  $\text{sign}(B_t)$  belongs to the Sobolev space  $\mathbb{D}^{\alpha,2}$  for any  $\alpha < \frac{1}{2}$ .

Now it is easy to obtain the chaos expansion of  $\int_0^t \text{sign}(B_s) dB_s$ .

**Proposition 5** For any  $0 < H < 1$ ,

$$\int_0^t \text{sign}(B_s) dB_s = \sum_{k=1}^{\infty} c_k I_{2k}(h_{2k}),$$

where

$$c_k = \frac{(-1)^{k-1}}{\sqrt{2\pi}(2k-1)(k-1)!2^{k-2}}$$

and

$$h_{2k}(s_1, \dots, s_{2k}) = (s_1 \vee s_2 \vee \dots \vee s_{2k})^{-(2k-1)H}.$$

A consequence of this proposition is the following

**Proposition 6** *For any  $0 < H < 1$  and  $t > 0$ , the random variable  $\int_0^t \text{sign}(B_s) dB_s$  belongs to the Sobolev space  $\mathbb{D}^{\alpha,2}$  for any  $\alpha < 1/2$ .*

**Proof.** It is easy to check that there is a constant  $C > 0$  such that

$$\mathbb{E} [I_{2k}(h_{2k})]^2 \leq C \frac{(2k)!}{(2k-1)^2 [(k-1)!]^2 2^{2k}}.$$

Therefore

$$\begin{aligned} \mathbb{E} [c_k I_{2k}(h_{2k})]^2 &\leq C \frac{(2k)!}{(k!2^k)^2} \\ &\leq C k^{-3/2}. \end{aligned}$$

This proves the proposition. ■

The next proposition states that  $\int_0^t \text{sign}(B_t) dB_t$  is not a fractional Brownian motion.

**Proposition 7** *The process  $X = \{X_t, t \geq 0\}$  is not a fractional Brownian motion.*

**Proof.** Suppose that  $X$  is a fractional Brownian motion. Then it is a fractional Brownian motion with Hurst parameter  $H$  since it is self-similar with parameter  $H$ . Then, the process

$$Y_t = \int_0^t \eta_H(t, r) dX_r$$

must be a standard Brownian motion with respect to the filtration generated by  $X$ , where

$$\eta_H(t, r) = (K_H^*)^{-1} (\mathbf{1}_{[0,t]})(r).$$

We claim that

$$Y_t = Z_t, \quad (8)$$

where

$$Z_t = \int_0^t \eta_H(t, r) \text{sign}(B_r) dB_r. \quad (9)$$

In fact, set  $t_k^n = \frac{tk}{n}$ ,  $k = 0, \dots, n$ , and consider the approximations

$$Y_t^n = \sum_{k=1}^n \eta_H(t, t_{k-1}^n) (X_{t_k^n} - X_{t_{k-1}^n}).$$

We know that  $Y_t^n$  converges in  $L^2(\Omega)$  to  $Y_t$ , because the functions

$$\sum_{k=1}^n \eta_H(t, t_{k-1}^n) \mathbf{1}_{[t_{k-1}^n, t_k^n]}(r)$$

converge to  $\eta_H(t, r) \mathbf{1}_{[0,t]}(r)$  in the norm of the Hilbert space  $\mathcal{H}$ . On the other hand, by Definition 1, for any smooth and cylindrical random variable  $F \in \mathcal{S}_{\mathcal{H}}$  we have

$$\begin{aligned} \mathbb{E}(FY_t^n) &= \mathbb{E} \left( \left\langle D_r F, \sum_{k=1}^n \eta_H(t, t_{k-1}^n) \mathbf{1}_{[t_{k-1}^n, t_k^n]}(r) \text{sign}(B_r) \right\rangle_{\mathcal{H}} \right) \\ &= \mathbb{E} \left( \left\langle \Gamma_{H,T}^{*,a} \Gamma_{H,T}^* D_r F, \sum_{k=1}^n \eta_H(t, t_{k-1}^n) \mathbf{1}_{[t_{k-1}^n, t_k^n]}(r) \text{sign}(B_r) \right\rangle_{L^2(0,T)} \right). \end{aligned}$$

As before this converges to

$$\begin{aligned} &\mathbb{E} \left( \left\langle K^{*,a} K^* D_r F, \eta_H(t, r) \mathbf{1}_{[0,t]}(r) \text{sign}(B_r) \right\rangle_{L^2(0,T)} \right) \\ &= \mathbb{E}(FZ_t), \end{aligned}$$

as  $n$  tends to infinity. So (8) holds.

We can write, using Lemma 3

$$Z_t = \sum_{k=0}^{\infty} b_k \int_0^t \eta_H(t, r) r^{-(2k+1)H} I_{2k+1}(\mathbf{1}_{[0,r]}^{\otimes(2k+1)}) dB_r,$$

where

$$b_k = \frac{(-1)^k}{\sqrt{2\pi}(2k+1)k!2^{k-1}}. \quad (10)$$

So,

$$Z_t = \sum_{k=0}^{\infty} b_k I_{2k+2,t}(f_{2k+2}),$$

where

$$\begin{aligned} & f_{2k+2}(t, s_1, \dots, s_{2k+2}) \\ = & \text{symm} \left( \eta_H(t, s_{2k+2}) s_{2k+2}^{-(2k+1)H} \mathbf{1}_{[0,2k+2]}(s_1) \cdots \mathbf{1}_{[0,2k+2]}(s_{2k+1}) \right) \\ = & \frac{1}{2k+1} \eta_H(t, s_1 \vee \cdots \vee s_{2k+2}) (s_1 \vee \cdots \vee s_{2k+2})^{-(2k+1)H}, \end{aligned}$$

and  $I_{2k+2,t}(f)$  denotes  $I_{2k+2}(f \mathbf{1}_{[0,t]}^{\otimes(2k+2)})$ . We can transform these multiple stochastic integrals into integrals with respect to a standard Brownian motion, using the operator  $K_H^*$ . In this way we obtain

$$I_{2k+2,t}(f_{2k+2}) = I_{2k+2,t}^W(K_H^{*\otimes(2k+2)} f_{2k+2}),$$

and the process

$$Z_t = \sum_{k=0}^{\infty} b_k I_{2k+2,t}^W(K_H^{*\otimes(2k+2)} f_{2k+2})$$

is a Brownian motion with respect to the filtration generated by  $W$ . Hence, every component of the chaos expansion is a martingale with respect to the filtration generated by  $W$ . In particular, this implies that the coefficient of the second chaos  $K_H^{*\otimes 2} f_2(t, s_1, s_2)$  must not depend on  $t$ .

For  $H > \frac{1}{2}$  we have

$$\begin{aligned} K_H^{*\otimes 2} f_2(t, s_1, s_2) &= d_H^2 (s_1 s_2)^{\frac{1}{2}-H} \\ &\times \left[ I_{t-}^{(H-\frac{1}{2})\otimes 2} (u_1 u_2)^{-\frac{1}{2}} \eta_H(t, u_1 \vee u_2) \right] (s_1, s_2) \end{aligned}$$

where  $d_H = c_H \Gamma(H - \frac{1}{2})$ . We have used the fact that

$$(I_{t-}^\alpha f) \mathbf{1}_{[0,t]} = I_{T-}^\alpha (f \mathbf{1}_{[0,t]}) .$$

Then

$$\begin{aligned} & \left[ I_{t-}^{(H-\frac{1}{2})\otimes 2} (u_1 u_2)^{-\frac{1}{2}} \eta_H(t, u_1 \vee u_2) \right] (s_1, s_2) \\ = & \frac{1}{\Gamma(H - \frac{1}{2})^2} \int_{s_2}^t \int_{s_1}^t ((u_1 - s_1)(u_2 - s_2))^{-\frac{1}{2}} \\ & \times \eta_H(t, (u_1 - s_1) \vee (u_2 - s_2)) du_1 du_2. \end{aligned}$$

Taking  $t = \max(s_1, s_2)$ , we would have  $K_H^{*\otimes 2} f_2(t, s_1, s_2) = 0$ , because

$$\eta_H(t, r) \leq C t^{H-\frac{1}{2}} r^{H-\frac{1}{2}} (t-r)^{\frac{1}{2}-H}.$$

Hence,  $\eta_H(t, u_1 \vee u_2) = 0$ , which leads to a contradiction.

Suppose now that  $H < \frac{1}{2}$ . In this case we have

$$\begin{aligned} K_H^{*\otimes 2} f_2(t, s_1, s_2) &= e_H^2 (s_1 s_2)^{\frac{1}{2}-H} \\ &\quad \times \left[ D_{t-}^{(\frac{1}{2}-H)\otimes 2} (u_1 u_2)^{-\frac{1}{2}} \eta_H(t, u_1 \vee u_2) \right] (s_1, s_2), \end{aligned}$$

where  $e_H = c_H \Gamma(H + \frac{1}{2})$ , using again that  $(D_{t-}^\alpha f) \mathbf{1}_{[0,t]} = D_{T-}^\alpha (f \mathbf{1}_{[0,t]})$ . Notice that

$$\eta_H(t, r) = \frac{1}{e_H \Gamma(\frac{1}{2}-H)} r^{\frac{1}{2}-H} \int_r^t (y-r)^{-\frac{1}{2}-H} y^{H-\frac{1}{2}} dy.$$

As a consequence,  $\eta_H(t, r)$  behaves as  $C r^{\frac{1}{2}-H} (t-r)^{\frac{1}{2}-H}$ . We have

$$\begin{aligned} &\left[ D_{t-}^{(\frac{1}{2}-H)\otimes 2} (u_1 u_2)^{-\frac{1}{2}} \eta_H(t, u_1 \vee u_2) \right] (s_1, s_2) \\ &= \frac{1}{\Gamma(H + \frac{1}{2})^2} D_{t-}^{\frac{1}{2}-H} \left( \frac{(s_1 s_2)^{-\frac{1}{2}} \eta_H(t, s_1 \vee s_2)}{(t-s_1)^{\frac{1}{2}-H} (t-s_2)^{\frac{1}{2}-H}} \right. \\ &\quad + \left( \frac{1}{2} - H \right) \int_{s_1}^t \frac{(s_1 s_2)^{-\frac{1}{2}} \eta_H(t, s_1 \vee s_2) - (y s_2)^{-\frac{1}{2}} \eta_H(t, y \vee s_2)}{(y-s_1)^{\frac{3}{2}-H} (t-s_2)^{\frac{1}{2}-H}} dy \\ &\quad + \left( \frac{1}{2} - H \right) \int_{s_2}^t \frac{(s_1 s_2)^{-\frac{1}{2}} \eta_H(t, s_1 \vee s_2) - (y s_1)^{-\frac{1}{2}} \eta_H(t, y \vee s_1)}{(y-s_2)^{\frac{3}{2}-H} (t-s_1)^{\frac{1}{2}-H}} dy \\ &\quad + \left( \frac{1}{2} - H \right)^2 \int_{s_1}^t \int_{s_2}^t (y-s_1)^{H-\frac{3}{2}} (z-s_2)^{H-\frac{3}{2}} (s_1 s_2)^{-\frac{1}{2}} [\eta_H(t, s_1 \vee s_2) \right. \\ &\quad \left. - (y s_2)^{-\frac{1}{2}} \eta_H(t, y \vee s_2) - (z s_1)^{-\frac{1}{2}} \eta_H(t, z \vee s_1) + (y z)^{-\frac{1}{2}} \eta_H(t, y \vee z)] dy dz \right). \end{aligned}$$

Taking again  $t = \max(s_1, s_2)$ , we would have  $K_H^{*\otimes 2} f_2(t, s_1, s_2) = 0$  which leads to a contradiction. ■

Consider the covariance between two increments of the process  $X$ :

$$r(n) := \mathbb{E} [(X_{a+1} - X_a) (X_{n+1} - X_n)],$$

where  $0 < a \leq n$ . We say that  $X$  is long-range dependent if for any  $a > 0$ ,

$$\sum_{n \geq a} |r(n)| = \infty.$$

The next proposition studies the long-range dependence properties of the process  $X$ . We see that this property differs from that of fractional Brownian motion.

**Proposition 8** *Let  $X_t = \int_0^t \text{sign}(B_s) dB_s$ . If  $H \geq 2/3$ , then  $X_t$  is long-range dependent and if  $1/2 < H < 2/3$ , then  $X_t$  is not long-range dependent.*

**Proof.** From Lemma 3 we can deduce the Wiener chaos expansion of the random variable  $X_t$ . In fact, we have

$$X_t = \sum_{k=0}^{\infty} b_k I_{2k+2,t}(h_{2k+2}),$$

where  $b_k$  is defined in (10) and

$$h_{2k+2}(s_1, \dots, s_{2k+2}) = (s_1 \vee \dots \vee s_{2k+2})^{-(2k+1)H}.$$

Let us compute the covariance of  $X_s$  and  $X_t - X_r$ , where  $0 < r < t$ . From the Itô isometry of multiple stochastic integrals it follows that

$$\mathbb{E}[X_s(X_t - X_r)] = \sum_{k=0}^{\infty} b_k^2 \mathbb{E}[I_{2k+2,s}(h_{2k+2}) [I_{2k+2,t}(h_{2k+2}) - I_{2k+2,r}(h_{2k+2})]].$$

We have

$$\begin{aligned} \mathbb{E}[X_s(X_t - X_r)] &= \sum_{k=0}^{\infty} b_k^2 (2k+2)! \left\langle h_{2k+2} \mathbf{1}_{[0,s]}^{\otimes(k+2)}, h_{2k+2} \left( \mathbf{1}_{[0,t]}^{\otimes(k+2)} - \mathbf{1}_{[0,r]}^{\otimes(k+2)} \right) \right\rangle_{\mathcal{H}^{2k+2}} \\ &\geq 2b_0^2 \left\langle h_2 \mathbf{1}_{[0,s]}^{\otimes 2}, h_2 \left( \mathbf{1}_{[0,t]}^{\otimes 2} - \mathbf{1}_{[0,r]}^{\otimes 2} \right) \right\rangle_{\mathcal{H}^2} \\ &\geq \frac{b_0^2}{2} \int_r^t \int_0^r \int_0^s \int_0^s s_2^{-H} t_2^{-H} \phi(s_1, t_1) \phi(s_2, t_2) ds_1 ds_2 dt_1 dt_2, \end{aligned}$$

where  $\phi(s, t) = H(2H-1)|t-s|^{2H-2}$ .

Thus let  $s = 1$ ,  $r = n$ , and  $t = n + 1$  and we have

$$\begin{aligned}
r(n) &:= \mathbb{E}[(X_{a+1} - X_a)(X_{n+1} - X_n)] \\
&\geq C \int_n^{n+1} \int_0^n \int_a^{a+1} \int_a^{a+1} s_2^{-H} t_2^{-H} (t_1 - s_1)^{2H-2} (t_2 - s_2)^{2H-2} ds_1 ds_2 dt_1 dt_2 \\
&\geq C \int_n^{n+1} \int_{a+2}^n \int_a^{a+1} \int_a^{a+1} s_2^{-H} t_2^{-H} (t_1 - a - 1)^{2H-2} (t_2 - a - 1)^{2H-2} ds_1 ds_2 dt_1 dt_2 \\
&\approx Cn^{3H-3},
\end{aligned}$$

as  $n$  tends to infinity. Thus if  $H \geq 2/3$ ,  $\sum_{n \geq a} r(n) = \infty$ .

If  $H < 2/3$ , then we use another approach. Set, as before

$$r(n) = \mathbb{E} \left[ \left( \int_a^{a+1} \text{sign} B_t dB_t \right) \left( \int_n^{n+1} \text{sign} B_t dB_t \right) \right].$$

Using the formula for the expectation of the product of two divergence integrals we obtain

$$\begin{aligned}
r(n) &= \alpha_H \int_a^{a+1} \int_n^{n+1} \mathbb{E}(\text{sign} B_s \text{sign} B_t) |s - t|^{2H-2} ds dt \\
&\quad + 4\alpha_H^2 \int_a^{a+1} \int_n^{n+1} \int_0^s \int_0^t \mathbb{E}(\delta_0(B_s) \delta_0(B_t)) \\
&\quad \times |s - \sigma|^{2H-2} |\theta - t|^{2H-2} d\sigma d\theta ds dt,
\end{aligned}$$

where  $\alpha_H = H(2H - 1)$ . This formula can be proved by approximating the function  $\text{sign}(x)$  by smooth functions and then taking the limit in  $L^2(\Omega)$ . We have

$$\int_0^t |s - \sigma|^{2H-2} d\sigma = \frac{1}{2H-1} (s^{2H-1} + |s - t|^{2H-1} \text{sign}(t - s)).$$

Hence,

$$\begin{aligned}
r(n) &= \alpha_H \int_a^{a+1} \int_n^{n+1} \mathbb{E}(\text{sign} B_s \text{sign} B_t) |s - t|^{2H-2} ds dt \\
&\quad + 4H^2 \int_a^{a+1} \int_n^{n+1} \mathbb{E}(\delta_0(B_s) \delta_0(B_t)) \\
&\quad \times (s^{2H-1} + (t - s)^{2H-1})(t^{2H-1} - (t - s)^{2H-1}) ds dt \\
&= a_n + b_n.
\end{aligned}$$

For the second term we have

$$b_n = \frac{4H^2}{2\pi} \int_a^{a+1} \int_n^{n+1} \frac{(s^{2H-1} + (t-s)^{2H-1})(t^{2H-1} - (t-s)^{2H-1})}{[(st)^{2H} - \frac{1}{4}(t^{2H} + s^{2H} - |t-s|^{2H})^2]^{1/2}} ds dt$$

Therefore,

$$\begin{aligned} b_n &\leq \frac{4H^2(2H-1)}{2\pi} \frac{((a+1)^{2H-1} + (n+1)^{2H-1})(n-a-1)^{2H-2}}{[(an)^{2H} - \frac{1}{4}((n+1)^{2H} + (a+1)^{2H} - |n-a|^{2H})^2]^{1/2}} \\ &\leq Cn^{3H-3}. \end{aligned}$$

To estimate  $a_n$ , we have from (6)

$$\begin{aligned} \mathbb{E}[\text{sign}(B_u)\text{sign}(B_v)] &= \sum_{k=0}^{\infty} \frac{4(2k)!}{(2k+1)^2 2\pi (k!2^k)^2 (uv)^{(2k+1)H}} (u^{2H} + v^{2H} - |u-v|^{2H})^{2k+1} \\ &\leq C \frac{u^{2H} + v^{2H} - |u-v|^{2H}}{(uv)^H} \sum_{k=0}^{\infty} k^{-3/2} \frac{(u^{2H} + v^{2H} - |u-v|^{2H})^{2k}}{(uv)^{2kH}} \\ &\leq C_1 \frac{u^{2H} + v^{2H} - |u-v|^{2H}}{(uv)^H}. \end{aligned}$$

Therefore

$$\begin{aligned} a_n &\leq C_2 \int_a^{a+1} \int_n^{n+1} |u-v|^{2H-2} \frac{u^{2H} + v^{2H} - |u-v|^{2H}}{(uv)^H} dv du \\ &\leq C_3 n^{3H-3}. \end{aligned}$$

As a consequence, if  $H < 2/3$ , then  $\sum_{n \geq 1} r(n) < \infty$ . ■

## 4 General Dimension

In this section we consider a  $d$ -dimensional fractional Brownian motion  $B = \{(B_t^1, \dots, B_t^d), t \geq 0\}$ , with Hurst parameter  $H > \frac{1}{2}$ . Let  $R_t = |B_t|$  be the fractional Bessel process associated to the  $d$ -dimensional fBm  $B$ .

Suppose first that  $H > \frac{1}{2}$ . Fix a time interval  $[0, T]$ , and define the derivative and divergence operators,  $D^{(i)}$  and  $\delta^{(i)}$ , with respect to each component

$B^{(i)}$ , as in Section 2. We assume that the Sobolev spaces  $\mathbb{D}_i^{1,p}$  include functionals of all the components of  $B$  and not only of component  $i$ . For each  $p > 1$ , let  $\mathbb{L}_{H,i}^{1,p}$  be the set of processes  $u \in \mathbb{D}_i^{1,p}(\mathcal{H})$  such that

$$\mathbb{E} \left[ \|u\|_{L^{1/H}([0,T])}^p \right] + \mathbb{E} \left[ \|D^{(i)}u\|_{L^{1/H}([0,T]^2)}^p \right] < \infty.$$

It has been proved in [1] that  $\left\{ \frac{B_s^i}{R_s}, s \in [0, T] \right\}$  belongs to the space  $\mathbb{L}_{H,i}^{1,1/H}$  for each  $i = 1, \dots, d$  and

$$R_t = \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} dB_s^i + H(d-1) \int_0^t \frac{s^{2H-1}}{R_s} ds. \quad (11)$$

In the case  $H < \frac{1}{2}$  the following result holds.

**Proposition 9** *If  $H < \frac{1}{2}$ , the process  $\frac{B_s^i}{R_s}$  belongs to the extended domain of the divergence operator  $\text{Dom}_t^* \delta^i$  on any time interval  $[0, t]$ , and (11) holds.*

**Proof.** For any test random variable  $F \in \mathcal{S}_{\mathcal{H}}$  we have

$$\begin{aligned} & \int_0^t \mathbb{E} \left( \frac{B_s^i}{R_s} K_H^{*,a} K_H^* D_s^{(i)} F \right) ds \\ &= \lim_{\varepsilon \downarrow 0} \int_0^t \mathbb{E} \left( h'_\varepsilon(R_s^2) B_s^i K_H^{*,a} K_H^* D_s^{(i)} F \right) ds \\ &= \lim_{\varepsilon \downarrow 0} \mathbb{E} \left( F \int_0^t h'_\varepsilon(R_s^2) B_s^i dB_s^i \right), \end{aligned}$$

where, for any  $\varepsilon > 0$ ,

$$h_\varepsilon(x) = \begin{cases} \frac{3}{8}\sqrt{\varepsilon} + \frac{3}{4\sqrt{\varepsilon}}x - \frac{1}{8\varepsilon\sqrt{\varepsilon}}x^2 & \text{if } x < \varepsilon \\ \sqrt{x} & \text{if } x \geq \varepsilon \end{cases}.$$

We have  $h_\varepsilon(x) \in C^2(\mathbb{R})$  and  $\lim_{\varepsilon \rightarrow 0} h_\varepsilon(x) = \sqrt{x}$  for all  $x \geq 0$ . By Itô's formula for the fractional Brownian motion in the case  $H < \frac{1}{2}$ , we have

$$h_\varepsilon(R_t^2) - h_\varepsilon(0) = \sum_{i=1}^d \int_0^t h'_\varepsilon(R_s^2) B_s^i dB_s^i + J_\varepsilon,$$

where

$$\begin{aligned} J_\varepsilon &= H(d-1) \int_0^t \mathbf{1}_{\{R_s^2 \geq \varepsilon\}} \frac{s^{2H-1}}{R_s} ds \\ &\quad + H \int_0^t \mathbf{1}_{\{R_s^2 < \varepsilon\}} \frac{1}{2\sqrt{\varepsilon}} \left[ 3d - (d+2) \frac{R_s^2}{\varepsilon} \right] s^{2H-1} ds. \end{aligned}$$

■

The process

$$X_t = \sum_{i=1}^d \int_0^t \frac{B_s^i}{R_s} \delta B_s^i \quad (12)$$

is  $H$ -self-similar, Hölder continuous of order  $\alpha < H$ , and it has the same  $\frac{1}{H}$ -variation as the fractional Brownian motion. Nevertheless, as we will show in the next proposition it is not a fractional Brownian motion with Hurst parameter  $H$ .

For  $h \in \mathcal{H}^{\otimes n}$ , we denote

$$I_{j_1, \dots, j_n}(h) = \int_{0 < s_1, \dots, s_n < t} h(s_1, \dots, s_n) dB_{s_1}^{j_1} \cdots dB_{s_n}^{j_n}, \quad (13)$$

First we find the chaos expansion of  $\sum_{i=1}^d \int_0^t f_i(B_s) dB_s^i$ , where  $f_i : \mathbb{R}^d \rightarrow \mathbb{R}$  are smooth functions with polynomial growth.

**Proposition 10** *The following chaos expansion holds for  $Z_t = \sum_{i=1}^d \int_0^t f_i(B_s) dB_s^i$*

$$Z_t = \sum_{i=1}^d \sum_{n=1}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} I_{j_1, \dots, j_n, i} (g_{j_1, \dots, j_n}^i(s_1, \dots, s_{n+1})),$$

where

$$\begin{aligned} g_{j_1, \dots, j_n}^i(s_1, \dots, s_{n+1}) &= \frac{(-1)^n (s_1 \vee \cdots \vee s_{n+1})^{-nH}}{(2\pi)^{d/2}} \\ &\quad \times \int_{\mathbb{R}^d} \left[ \frac{\partial^n}{\partial y_{j_1} \cdots \partial y_{j_n}} e^{-\frac{|y|^2}{2}} \right] f_i(y(s_1 \vee \cdots \vee s_{n+1})^H) dy. \end{aligned}$$

**Proof.** Using Stroock's formula yields for each  $i = 1, \dots, d$

$$f_i(B_s) = \sum_{n=0}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} \frac{1}{n!} I_{j_1, \dots, j_n} \left( f_{j_1, \dots, j_n}^i(s) \mathbf{1}_{[0,s]}^{\otimes n} \right),$$

where

$$\begin{aligned}
f_{j_1, \dots, j_n}^i(s) &= \mathbb{E}(D^{j_1} \cdots D^{j_n}(f_i(B_s))) \\
&= \mathbb{E}\left(\frac{\partial^n f_i}{\partial z_{j_1} \cdots \partial z_{j_n}}(B_s)\right) \\
&= \frac{1}{(2\pi s^{2H})^{d/2}} \int_{\mathbb{R}^d} \frac{\partial^n f_i}{\partial z_{j_1} \cdots \partial z_{j_n}}(z) e^{-\frac{|z|^2}{2s^{2H}}} dz \\
&= \frac{s^{-nH}}{(2\pi)^{d/2}} \int_{\mathbb{R}^d} \frac{\partial^n f_i}{\partial y_{j_1} \cdots \partial y_{j_n}}(ys^H) e^{-\frac{|y|^2}{2y}} dy \\
&= \frac{(-1)^n s^{-nH}}{(2\pi)^{d/2}} \int_{\mathbb{R}^d} f_i(ys^H) \frac{\partial^n}{\partial y_{j_1} \cdots \partial y_{j_n}} e^{-\frac{|y|^2}{2y}} dy.
\end{aligned}$$

Finally, the result follows from

$$\begin{aligned}
Z_t &= \sum_{i=1}^d \int_0^t f_i(B_s) dB_s^i \\
&= \sum_{i=1}^d \sum_{n=0}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} I_{j_1, \dots, j_n, i} \left( \text{symm} \left( f_{j_1, \dots, j_n}^i(s) \mathbf{1}_{[0,s]}^{\otimes n} \mathbf{1}_{[0,t]}(s) \right) \right) \\
&= \sum_{i=1}^d \sum_{n=0}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} I_{j_1, \dots, j_n, i} \left( f_{j_1, \dots, j_n}^i(s_1 \vee \cdots \vee s_{n+1}) \prod_{i=1}^{n+1} \mathbf{1}_{[0,t]}(s_i) \right),
\end{aligned}$$

which completes the proof of the proposition. ■

Now let  $f_i(x) = \frac{x_i}{\sqrt{x_1^2 + \cdots + x_d^2}}$ . Then it is easy to check  $f_i(tx) = f_i(x)$  for all  $t > 0$ . Hence, for such  $f_i$ , we have

$$g_{j_1, \dots, j_n}^i(s_1, \dots, s_{n+1}) = b_{j_1, \dots, j_n}(s_1 \vee \cdots \vee s_{n+1})^{-nH},$$

where

$$b_{i, j_1, \dots, j_n} = \frac{(-1)^n}{(2\pi)^{d/2}} \int_{\mathbb{R}^d} \left[ \frac{\partial^n}{\partial y_{j_1} \cdots \partial y_{j_n}} e^{-\frac{|y|^2}{2}} \right] f_i(y) dy.$$

Then the chaos expansion of  $f_i(B_t)$  is given by

$$f_i(B_t) = \sum_{n=0}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} b_{i, j_1, \dots, j_n} t^{-nH} \int_{\{0 < s_1 < \cdots < s_n < t\}} dB_{s_1}^{j_1} \cdots dB_{s_n}^{j_n},$$

and the chaos expansion of the divergence of this process is

$$\begin{aligned} \int_0^t f_i(B_s) dB_s^i &= \sum_{n=1}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} b_{i,j_1, \dots, j_n} \\ &\quad \times \int_{\{0 < s_1, \dots, s_{n+1} < t\}} (s_1 \vee \dots \vee s_{n+1})^{-nH} dB_{s_1}^{j_1} \cdots dB_{s_n}^{j_n} dB_{s_{n+1}}^i. \end{aligned}$$

Using these results we can prove the following proposition.

**Proposition 11** *The process  $Y = \{Y_t, t \geq 0\}$  defined in (12) is not a fractional Brownian motion.*

**Proof.** If  $Y_t = \sum_{i=1}^d \int_0^t f_i(B_s) dB_s^i$  is a fractional Brownian motion, then as in previous section one can show that

$$Z_t = \sum_{i=1}^d \int_0^t \eta_H(t, s) f_i(B_s) dB_s^i$$

is a classical Brownian motion. But

$$Z_t = \sum_{i=1}^d \sum_{n=1}^{\infty} \sum_{1 \leq j_1, \dots, j_n \leq d} b_{i,j_1, \dots, j_n} \int_{\{0 < s_1, \dots, s_{n+1} < t\}} h(t, s_1, \dots, s_{n+1}) dB_{s_1}^{j_1} \cdots dB_{s_n}^{j_n} dB_{s_{n+1}}^i,$$

where

$$h(t, s_1, \dots, s_{n+1}) = \eta(t, s_1 \vee s_2 \vee \dots \vee s_{n+1}) (s_1 \vee s_2 \vee \dots \vee s_{n+1})^{-nH}.$$

In a similar way to one dimensional case, one can show that  $\{Z_t, t \geq 0\}$  is not a martingale ■

**Proposition 12** *Let  $Y = \{Y_t, t \geq 0\}$  be the process defined in (12). If  $H \geq 2/3$ , then  $Y_t$  is long range dependent and if  $1/2 < H < 2/3$ , then  $Y_t$  is not long range dependent.*

**Proof.** Set

$$\rho_n = \mathbb{E} \left[ \left( \sum_{i=1}^d \int_0^1 \frac{B_s^i}{|B_s|} \delta B_s^i \right) \left( \sum_{i=1}^d \int_n^{n+1} \frac{B_s^i}{|B_s|} \delta B_s^i \right) \right].$$

By the formula for the expectation of the product of two divergence integrals we can write

$$\begin{aligned}
\rho_n &= \sum_{i,j=1}^d \alpha_H \int_0^1 \int_n^{n+1} \mathbb{E} \left( \frac{B_s^i B_t^i}{|B_s| |B_t|} \right) |t-s|^{2H-2} ds dt \\
&\quad + \sum_{i,j=1}^d \alpha_H^2 \int_0^1 \int_n^{n+1} \int_0^t \int_0^s \mathbb{E} \left( D_\theta^j \left( \frac{B_s^i}{|B_s|} \right) D_\sigma^i \left( \frac{B_t^j}{|B_t|} \right) \right) \\
&\quad \times |\theta-t|^{2H-2} |\sigma-s|^{2H-2} d\theta d\sigma ds dt \\
&:= \rho_n^1 + \rho_n^2.
\end{aligned}$$

In order to estimate the term  $\rho_n^1$  we make use of the orthogonal decomposition

$$B_t = \frac{R(t, s)}{s^{2H}} B_s + \beta_{s,t} Y,$$

where

$$\beta_{s,t}^2 = \frac{(st)^{2H} - R(t, s)^2}{s^{2H}},$$

and  $Y$  is a  $d$ -dimensional standard normal random variable independent of  $B_s$ . Set

$$\lambda_{st} = \frac{R(t, s)}{\beta_{s,t} s^{2H}} = \frac{\frac{1}{2} (t^{2H} + s^{2H} - |t-s|^{2H})}{s^H [(st)^{2H} - \frac{1}{4} (t^{2H} + s^{2H} - |t-s|^{2H})^2]^{1/2}}$$

As  $t$  tends to infinity and  $s$  belongs to  $(0, 1)$ , the term  $\lambda_{st}$  behaves as  $H s^{-2H} t^{H-1}$ . Hence, by Lemma 13

$$\begin{aligned}
\mathbb{E} \left( \frac{\langle B_s, B_t \rangle}{|B_s| |B_t|} \right) &= \mathbb{E} \left( \frac{\langle B_s, \lambda_{st} B_s + Y \rangle}{|B_s| |\lambda_{st} B_s + Y|} \right) \\
&= \mathbb{E} \left( \frac{\langle B_1, s^H \lambda_{st} B_1 + Y \rangle}{|B_1| |s^H \lambda_{st} B_1 + Y|} \right) \\
&= s^H \lambda_{st} \mathbb{E} \left( \frac{|B_1|^2 |Y|^2 - \langle B_1, Y \rangle^2}{|B_1| |Y|^3} \right) + o(t^{H-1}).
\end{aligned}$$

Hence,

$$\mathbb{E} \left( \frac{\langle B_s, B_t \rangle}{|B_s| |B_t|} \right) \approx H C s^{-H} t^{H-1},$$

where

$$C = \mathbb{E} \left( \frac{|B_1|^2 |Y|^2 - \langle B_1, Y \rangle^2}{|B_1| |Y|^3} \right) > 0.$$

This implies that the term  $\rho_n^1$  behaves as  $n^{3H-3}$ .

For the term  $\rho_n^2$  we have

$$\begin{aligned} \rho_n^2 &= H^2 \sum_{i,j=1}^d \int_0^1 \int_n^{n+1} \mathbb{E} \left( \left( \frac{\delta_{ij}}{|B_s|} - \frac{B_s^i B_s^j}{|B_s|^3} \right) \left( \frac{\delta_{ij}}{|B_t|} - \frac{B_t^i B_t^j}{|B_t|^3} \right) \right) \\ &\quad \times (s^{2H-1} + (t-s)^{2H-1})(t^{2H-1} - (t-s)^{2H-1}) ds dt \\ &= H^2 \int_0^1 \int_n^{n+1} \mathbb{E} \left( \frac{d}{|B_s| |B_t|} - \frac{|B_t|^2}{|B_s| |B_t|^3} - \frac{|B_s|^2}{|B_s|^3 |B_t|} + \frac{\langle B_s, B_t \rangle^2}{|B_s|^3 |B_t|^3} \right) \\ &\quad \times (s^{2H-1} + (t-s)^{2H-1})(t^{2H-1} - (t-s)^{2H-1}) ds dt \\ &= H^2 \int_0^1 \int_n^{n+1} \mathbb{E} \left( \frac{d-2}{|B_s| |B_t|} + \frac{\langle B_s, B_t \rangle^2}{|B_s|^3 |B_t|^3} \right) \\ &\quad \times (s^{2H-1} + (t-s)^{2H-1})(t^{2H-1} - (t-s)^{2H-1}) ds dt. \end{aligned}$$

The term

$$\mathbb{E} \left( \frac{1}{|B_s| |B_t|} \left( d-2 + \frac{\langle B_s, B_t \rangle^2}{|B_s|^2 |B_t|^2} \right) \right)$$

behaves as  $Kt^{-H}$  as  $t$  tends to infinity, where

$$K = \mathbb{E} \left( \frac{1}{|B_1| |Y|} \left( d-2 + \frac{\langle B_1, Y \rangle^2}{|B_1|^2 |Y|^2} \right) \right) > 0.$$

Hence, the term  $\rho_n^2$  behaves also as  $n^{3H-3}$ . This completes the proof taking into account that the constants  $C$  and  $K$  are positive. ■

**Lemma 13** *Let  $X$  and  $Y$  be independent  $d$ -dimensional standard normal random variables. Then as  $\varepsilon$  tends to zero we have*

$$\mathbb{E} \left( \frac{\langle X, \varepsilon X + Y \rangle}{|X| |\varepsilon X + Y|} \right) = \varepsilon \mathbb{E} \left( \frac{|X|^2 |Y|^2 - \langle X, Y \rangle^2}{|X| |Y|^3} \right) + o(\varepsilon).$$

**Proof.** We have

$$\begin{aligned}
\mathbb{E} \left( \frac{\langle X, \varepsilon X + Y \rangle}{|X| |\varepsilon X + Y|} \right) &= \mathbb{E} \left( \frac{\langle X, \varepsilon X + Y \rangle}{|X| |\varepsilon X + Y|} - \frac{\langle X, Y \rangle}{|X| |Y|} \right) \\
&= \mathbb{E} \left( \frac{\langle X, \varepsilon X + Y \rangle |Y| - \langle X, Y \rangle |\varepsilon X + Y|}{|X| |\varepsilon X + Y| |Y|} \right) \\
&= \mathbb{E} \left( \frac{\varepsilon |X|^2 |Y| - \varepsilon \langle X, Y \rangle^2 / |Y| + o(\varepsilon)}{|X| |\varepsilon X + Y| |Y|} \right),
\end{aligned}$$

and that yields the desired estimation. ■

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